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First year Licence Introduction to probability and descriptive statistics

## Answers of the third series : Combinatorial analysis

Answer 01 : 1)Show that:

$$
C_{n}^{1}+C_{n}^{3}+\ldots=C_{n}^{0}+C_{n}^{2}+\ldots \quad \text { for any } n
$$

We have

$$
(x+y)^{n}=\sum_{i=0}^{n} C_{n}^{k} x^{n-k} y^{k}
$$

We pose $x=1$ and $y=-1$, and the previous equation reduces to :

$$
0=\sum_{i=0}^{n} C_{n}^{k}(-1)^{k}
$$

Wich can be written

$$
C_{n}^{0}+C_{n}^{2}+\ldots=C_{n}^{1}+C_{n}^{3}+\ldots
$$

2) Prove that

$$
C_{n}^{1}+2 C_{n}^{2}+\ldots+n C_{n}^{n}=n 2^{n-1}
$$

This time we begin with the expansion of $(1+x)^{n}$ :

$$
(1+x)^{n}=\sum_{i=0}^{n} C_{n}^{k} 1^{n-k} x^{k}
$$

Differentiating both sides of the previous equation with respect to $x$ gives

$$
n(1+x)^{n-1}=\sum_{i=1}^{n} C_{n}^{k} k x^{k-1}
$$

Now, let $x=1$, we find :

$$
n 2^{n-1}=\sum_{i=1}^{n} C_{n}^{k} k x^{k-1}=C_{n}^{1}+2 C_{n}^{2}+\ldots+n C_{n}^{n}
$$

Answer 03 : Roll 3 different dice simultaneously and by using the results obtained we construct a 3 -digit number.

1. $6 \times 6 \times 6=6^{3}$
2. The number of numbers are less than 500 and greater than 200 is: $3 \times 6^{2}$.
3. $3 \times 6^{2}$.
4. $6 \times 5 \times 4=P_{6}^{3}$

Answer 04 : Passwords have 3 different letters followed by 2 different symbols of the following set $\{@, \%, \$, \star\}$ then 2 numbers.
a) $26 \times 25 \times 24 \times 4 \times 3 \times 10 \times 10=P_{26}^{3} \times P_{4}^{2} \times 10^{2}$.
b) $6 \times 25 \times 24 \times 4 \times 3 \times 10 \times 5=6 \times P_{25}^{2} \times P_{4}^{2} \times 10 \times 5$.

Answer 05 : 1)a) For the word "Maths"? the number of arrangement is : 5!
b) For the word "proposition" we have $\{p, p, r, o, o, o, s, i, i, t, n\}$ so the number of arrangement is : $\frac{11!}{2!\times 3!\times 2 \text { ! }}$.
c) For the word "theorem" we have : $\frac{7!}{2!}$.
d) For the word "arrangement" we have : $\frac{11!}{2!\times 2!\times 2!\times 2!}$.

Answer 06 : Twenty books are to be arranged on a shelf; eleven on travel, five on cooking, and four on gardening.

1. 20 !.
2. We have three groups so 3 !. For the books on travel we have 11!, for the books on cooking we have 5 ! and for the books on gardening we have 4!. So the number of arrangements is : $3!\times 11!\times 5!\times 4$ !.
3. 11 ! $\times(5+4+1)$ ! or 11 ! $\times(20-11+1)$ !.

## Answer 07 :

1) There are $C_{70}^{12}$ choices for the first group. Having chosen 12 for the first group, there are $C_{58}^{12}$ choices for the second group and so on. The total is :

$$
C_{70}^{12} \times C_{58}^{12} \times C_{46}^{12} \times C_{34}^{12} \times C_{22}^{12} \times C_{10}^{10}=\frac{70!}{(12!)^{5} \times 10!}
$$

## answer of the fourth series : Probability space

Answer 01 : Express each of the following events by using the events $A, B$ and $C$ :

1. Exactly $A$ occurs : $A \cap \bar{B} \cap \bar{C}$.
2. $A$ and $B$ occur : $A \cap B \cap \bar{C}$.
3. All three events occur : $A \cap B \cap C$.
4. Exactly one of the three events occurs : $(A \cap \bar{B} \cap \bar{C}) \cup(\bar{A} \cap B \cap \bar{C}) \cup(\bar{A} \cap \bar{B} \cap C)$.
5. Exactly two of the three events occurs : $(A \cap B \cap \bar{C}) \cup(A \cap \bar{B} \cap C) \cup(\bar{A} \cap B \cap C)$..
6. None of the three events occurs : $(\bar{A} \cap \bar{B} \cap \bar{C})$.
7. At least one of the events occurs : $A \cup B \cup C$.
8. At most one of the events occurs : $4 \cup 6$.

## Answer 03 :

1. let $(\Omega, \mathcal{F}, P)$ be a probability space.
i) We want to prove that $\mathcal{G}=\{A \in \mathcal{F}, \quad P(A)=0$ ou $P(A)=1\}$ is a tribe on $\Omega$.

Cond.1. $\mathcal{G}$ is no vide because $\Omega \in \mathcal{F}$ and $P(\Omega)=1$, so $\Omega \in \mathcal{G}$.
Cond.2. Let $A \in \mathcal{G}$ be an event and we show that $\bar{A} \in \mathcal{G}$.
We have

$$
\begin{aligned}
A \in \mathcal{G} & \Longrightarrow A \in \mathcal{F} \text { and } P(A)=0 \text { or } P(A)=1 \\
& \Longrightarrow \bar{A} \in \mathcal{F} \text { and } P(\bar{A})=1-P(A)=1-0=1 \quad \text { or } \quad P(\bar{A})=0 \\
& \Longrightarrow \bar{A} \in \mathcal{F} \text { and } P(\bar{A})=0 \text { or } P(\bar{A})=1 \\
& \Longrightarrow \bar{A} \in \mathcal{G} .
\end{aligned}
$$

Cond.3. Let $\left(A_{i}\right)_{i \geq 0}$ be a sequence of events of $\mathcal{G}$. We want to prove that $\cup_{i} A_{i} \in \mathcal{G}$.

$$
\begin{aligned}
\text { For all } i \geq 0, A_{i} \in \mathcal{G} & \Longrightarrow \forall i \geq 0, A_{i} \in \mathcal{F} \text { and } P\left(A_{i}\right)=0 \text { or } P\left(A_{i}\right)=1 \\
& \Longrightarrow \cup_{i} A_{i} \in \mathcal{F} \ldots \ldots \ldots . .(1) \text { becaus } \mathcal{F} \text { is a tribe }
\end{aligned}
$$

If $P\left(A_{i}\right)=0$, for all $i \geq 0$, so $P\left(\cup_{i} A_{i}\right)=0$. $\qquad$
If at least $A_{i}, i \geq 0$, such that $P\left(A_{i}\right)=1$, we find $P\left(\cup_{i} A_{i}\right)=1$.
So from (1), (2), and (3) we deduce that $\cup_{i} A_{i} \in \mathcal{G}$.
ii) Let $A \in \mathcal{F}$ be an event. We want to show that

$$
\mathcal{F}_{A}=\{A \cap B, \quad B \in \mathcal{F}\}
$$

is a tribe on $\Omega$.

Cond.1. We have $A \in \mathcal{F}_{A}$ because $A \cap \Omega=A$ and $\Omega \in \mathcal{F}$. Or,
we have $\emptyset \in \mathcal{F}_{A}$ because $A \cap \emptyset$ and $\emptyset \in \mathcal{F}$.
Cond.2. Let $C \in \mathcal{F}_{A}$ be an avent and we want to prove that $\bar{C} \in \mathcal{F}_{A}$.
First, note that $\bar{C}$ is the complement of $C$ with respect to $A$, so we can write

$$
\bar{C}=A \backslash C=A \cap \bar{C}
$$

So we have

$$
\begin{aligned}
C \in \mathcal{F}_{A} & \Longrightarrow \exists B \in \mathcal{F}, C=A \cap B \\
& \Longrightarrow \exists B \in \mathcal{F}, \quad \bar{C}=A \cap(\overline{A \cap B}) \\
& \Longrightarrow \exists B \in \mathcal{F}, \quad \bar{C}=A \cap(\bar{A} \cup \bar{B}) \\
& \Longrightarrow \exists B \in \mathcal{F}, \quad \bar{C}=(A \cap \bar{A}) \cup(A \cap \bar{B}) \\
& \Longrightarrow \exists B \in \mathcal{F}, \quad \bar{C}=A \cap \bar{B} \text { et } \bar{B} \in \mathcal{F} \\
& \Longrightarrow \bar{C}=A \cap \bar{B} \in \mathcal{F}_{A}
\end{aligned}
$$

Cond.3. Let $\left(A_{i}\right)_{i \geq 0}$ be a sequence of events of $\mathcal{F}_{A}$.
We want to prove that $\cup_{i} A_{i} \in \mathcal{F}_{A}$. We have :

$$
\begin{aligned}
\text { For all } i \geq 0, A_{i} \in \mathcal{F}_{A} & \Longrightarrow \forall i \geq 0, \exists B_{i} \in \mathcal{F}, A_{i}=A \cap B_{i} \\
& \Longrightarrow \cup_{i} A_{i}=\cup_{i}\left(A \cap B_{i}\right) \\
& \Longrightarrow \cup_{i} A_{i}=A \cap\left(\cup_{i} B_{i}\right) \\
& \Longrightarrow \cup_{i} A_{i} \in \mathcal{F}_{A}, \text { because } \cup_{i} B_{i} \in \mathcal{F}
\end{aligned}
$$

2. $\star$ Let $\left(\Omega, \mathcal{F}_{i}\right), i \in I$, be a probable space. Prove that $\mathcal{F}=\cap_{i \in I} \mathcal{F}_{i}$ is a tribe on $\Omega$.

Answer 04 : Let $(\Omega, \mathcal{F}, P)$ be a probability space.

1) Let $B$ be an event. Show that $P_{B}$ is a probability on $\Omega$ such as:

$$
P_{B}(A)=\frac{P(A \cap B)}{P(B)}=P(A \mid B) .
$$

Cond.1. $0 \leq P_{B}(A) \leq 1$, because
$A \cap B \subseteq B \Rightarrow 0 \leq P(A \cap B) \leq P(B) \Rightarrow 0 \leq \frac{P(A \cap B)}{P(B)} \leq 1$.
Cond.2. $\quad P_{B}(\Omega)=1$, because $P_{B}(\Omega)=\frac{P(\Omega \cap B)}{P(B)}=\frac{P(B)}{P(B)}=1$
Cond.3. Let $\left(A_{i}\right)_{i \geq 0}$ be a sequence of events of $\mathcal{F}$, such as these events are disjoints two by two.
We want to prove that $P_{B}\left(\cup_{i} A_{i}\right)=\sum_{i} P_{B}\left(A_{i}\right)$. We have :

$$
\begin{aligned}
P_{B}\left(\cup_{i} A_{i}\right) & =\frac{P\left(\left(\cup_{i} A_{i}\right) \cap B\right)}{P(B)} \\
& =\frac{P\left(\cup_{i}\left(A_{i} \cap B\right)\right)}{P(B)}
\end{aligned}
$$

Note that the events $A_{i}, i=1, \ldots, n$, are disjoints two by two, then the events $A_{i} \cap B$ are also disjoints two by two. So :

$$
\begin{aligned}
P_{B}\left(\cup_{i} A_{i}\right) & =\sum_{i} \frac{P\left(A_{i} \cap B\right)}{P(B)} \\
& =\sum_{i} P_{B}\left(A_{i}\right)
\end{aligned}
$$

2) $\star$ Let $a$ be a reel and $\Omega=\mathbb{R}$. Show that $P_{a}$ is a probability on $\Omega$ such as :

$$
P_{a}(A)=\delta_{a}(A)=\left\{\begin{array}{lll}
1 & \text { if } & a \in A \\
0 & \text { if } & \text { non }
\end{array}\right.
$$

Cond.1. $0 \leq P_{a}(A) \leq 1$.
Cond.2. $\quad P_{a}(\Omega)=1$, because $a \in \mathbb{R}=\Omega$.
Cond.3. Let $\left(A_{i}\right)_{i \geq 0}$ be a sequence of events of $\mathcal{F}$ two by two disjoints.
We want to prove that $P_{a}\left(\cup_{i} A_{i}\right)=\sum_{i} P_{a}\left(A_{i}\right)$. We have :

$$
P_{a}\left(\cup_{i} A_{i}\right)=\left\{\begin{array}{lll}
1 & \text { si } & a \in \cup_{i} A_{i} \\
0 & \text { si } & \text { non }
\end{array}\right.
$$

Case 1: $P_{a}\left(\cup_{i} A_{i}\right)=1$, so :

$$
\begin{aligned}
a \in \cup_{i} A_{i} \Leftrightarrow & \exists!i, a \in A_{i}, \text { and for all } j \neq i, a \notin A_{j} \\
& \text { because } A_{i} \cap A_{j}=\emptyset, \forall i \neq j \\
\Leftrightarrow & \exists!i, P_{a}\left(A_{i}\right)=1, \text { et } \forall j \neq i, P_{a}\left(A_{j}\right)=0
\end{aligned}
$$

so

$$
\begin{aligned}
\sum_{i} P_{a}\left(A_{i}\right) & =P_{a}\left(A_{0}\right)+\ldots+P_{a}\left(A_{i-1}\right)+P_{a}\left(A_{i}\right)+P_{a}\left(A_{i+1)}+\ldots\right. \\
& =0+\ldots+0+1+0+\ldots=1
\end{aligned}
$$

Case 2: $P_{a}\left(\cup_{i} A_{i}\right)=0$, so :

$$
\begin{aligned}
a \notin \cup_{i} A_{i} & \Leftrightarrow \forall i \geq 0, a \notin A_{i} \\
& \Leftrightarrow \forall i \geq 0, P_{a}\left(A_{i}\right)=0
\end{aligned}
$$

so

$$
\begin{aligned}
\sum_{i} P_{a}\left(A_{i}\right) & =P_{a}\left(A_{0}\right)+P_{a}\left(A_{1}\right)+P_{a}\left(A_{2}\right)+\ldots \\
& =0+0+0+\ldots=0
\end{aligned}
$$

Answer 05 : Let $(\Omega, \mathcal{F}, P)$ be a probability space and $A_{1}, \ldots, A_{n} n$ events of $\mathcal{F}$. Prove that:

1) By using " la démonstration par récurrence ", we prove the following proposition $P(n)$ :

$$
P(n): P\left(\cup_{i=1}^{n} A_{i}\right) \leq \sum_{i=1}^{n} P\left(A_{i}\right)
$$

Step 1: We want to prove the proposition for $n=2$ (i.e, we show that $P(n=2)$ is true).
So, we have :

$$
P\left(A_{1} \cup A_{2}\right)=P\left(A_{1}\right)+P\left(A_{2}\right)-P\left(A_{1} \cap A_{2}\right) \leq P\left(A_{1}\right)+P\left(A_{2}\right)=\sum_{i=1}^{2} P\left(A_{i}\right)
$$

Step 2: We assume that $P(n)$ is true, and we prove that the $P(n+1)$ is true.

$$
P\left(\cup_{i=1}^{n+1} A_{i}\right)=P\left(\left(\cup_{i=1}^{n} A_{i}\right) \cup A_{n+1}\right)
$$

We pose $B=\cup_{i=1}^{n} A_{i}$, so :

$$
\begin{aligned}
P\left(B \cup A_{n+1}\right) & =P(B)+P\left(A_{n+1}\right)-P\left(B \cap A_{n+1}\right) \\
& \leq P(B)+P\left(A_{n+1}\right)=P\left(\cup_{i=1}^{n} A_{i}\right)+P\left(A_{n+1}\right) \quad \text { by using the } P(n) \\
& \leq \sum_{i=1}^{n} P\left(A_{i}\right)+P\left(A_{n+1}\right) \\
& =\sum_{i=1}^{n+1} P\left(A_{i}\right) .
\end{aligned}
$$

2) We have :

$$
P\left(\cap_{i=1}^{n} A_{i}\right)=P\left(\overline{\cup_{i=1}^{n} \overline{A_{i}}}\right)=1-P\left(\cup_{i=1}^{n} \overline{A_{i}}\right)
$$

From 1), we find :

$$
\begin{aligned}
& P\left(\cup_{i=1}^{n} \overline{A_{i}}\right) \leq \sum_{i=1}^{n} P\left(\overline{A_{i}}\right) \\
\Rightarrow & 1-P\left(\cup_{i=1}^{n} \overline{A_{i}}\right) \geq 1-\sum_{i=1}^{n} P\left(\overline{A_{i}}\right) \\
\Rightarrow & P\left(\cap_{i=1}^{n} A_{i}\right) \geq 1-\sum_{i=1}^{n} P\left(\overline{A_{i}}\right)
\end{aligned}
$$

Answer 06 : Let
$M$ the event " observe a marked ball" so we have $P(M)=3 / 10$,
$B$ the event" observe a black ball" so we have $P(B)=7 / 10$,
$R$ the event " observe a red ball " so we have $P(R)=1-P(B)=3 / 10$, and $P(R \cap M)=2 / 10$.

1. $P(R \cup M)=P(R)+P(M)-P(R \cap M)=(1-P(N))+P(M)-P(R \cap M)=4 / 10$.
2. $P(R \cap \bar{M})=P(R \backslash M)=P(R)-P(R \cap M)=$ $\qquad$

## 3. The $1^{\text {st }}$ method :

$$
P(B \cap \bar{M})=P(\bar{R} \cap \bar{M})=P(\overline{R \cup M})=1-P(R \cup M)=
$$

$\qquad$
The $2^{\text {nd }}$ method :

$$
P(B \cap \bar{M})=P(B \backslash M)=P(B)-P(B \cap M)
$$

Find $P(B \cap M)$ : note that $\{B, R\}$ is the partition of $\Omega$ so :
$P(M)=P(M \cap B)+P(M \cap R) \rightarrow P(B \cap M)=P(M)-P(M \cap R)=(3 / 10)-(2 / 10)=1 / 10$.

## Answer 08 :

1. Let $W_{n}$ the event " draw two white balls ", so :

$$
P\left(W_{n}\right)=\frac{\operatorname{Card}\left(W_{n}\right)}{\operatorname{card}(\Omega)}=\frac{C_{n}^{2}}{C_{n+8}^{2}}=\frac{n \times(n-1)}{(n+8) \times(n+7)}
$$

2. Let $R$ the event "draw two red balls " and $G$ the event "draw two green balls ", so :

$$
\begin{gathered}
P(n)=P\left(W_{n}\right)+P(R)+P(G)=\frac{n \times(n-1)}{(n+8) \times(n+7)}+\frac{\operatorname{card}(R)+\operatorname{card}(G)}{\operatorname{card}(\Omega)} \\
=\frac{n \times(n-1)}{(n+8) \times(n+7)}+\frac{C_{5}^{2}+C_{3}^{2}}{C_{n+8}^{2}}=\frac{n^{2}-n+26}{(n+8) \times(n+7)}
\end{gathered}
$$

$\lim _{n \rightarrow \infty} P(n)=1$. When $n \rightarrow \infty$, we can say that the event "both balls are the same color" is almost certain.

Answer 09 : A queue is formed randomly by n people. Two friends A and B are in this queue. So, the number of arrangements for $n$ people is : $n!$.

1. The probability that the two friends are located one behind the other is :

$$
\frac{(n-1) \times 2!\times(n-2)!}{n!}=\frac{(n-1)!\times 2!}{n!}=\frac{2}{n}
$$

The two friends A and B form a set, and the number of arrangements the $n-2$ people is $(n-2)$ ! possible. Since A can be placed in line with B behind him or A behind B, so there are $(n-1) \times 2$ ! possible cases.

Or,

$$
\frac{2!\times(n-2+1)!}{n!}=\frac{2}{n}
$$

2. The probability that the two friends are r places apart (separated by $r-1$ people) is :

$$
\frac{(n-r) \times 2!\times(n-2)!}{n!}=\frac{2 \times(n-r)}{n \times(n-1)}
$$

The possible cases of finding person A then $r-1$ people then person B or the opposite are : $(n-r) \times 2$ !.

## Answer 10 :

1. Note that the probability of each face appearing is proportional to the number written on it, so :

$$
p_{1}=1 \times p, \quad p_{2}=2 p, \quad p_{3}=3 p, \quad p_{4}=4 p, \quad p_{5}=5 p, \quad p_{6}=6 p
$$

And we have : $\quad \sum_{i=1}^{6} p_{i}=1 \Rightarrow p=\frac{1}{21}$. We deduce the following table :

| Face i | 1 | 2 | 3 | 4 | 5 | 6 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $p_{i}=P(i)$ | $\frac{1}{21}$ | $\frac{2}{21}$ | $\frac{3}{21}$ | $\frac{4}{21}$ | $\frac{5}{21}$ | $\frac{6}{21}$ |

2. Let $A$ the event "obtain an even number", $A=\{2,4,6\}$, so

$$
P(A)=P(\{2\} \cup\{4\} \cup\{6\})=P(\{2\})+P(\{4\})+P(\{6\})=p_{2}+p_{4}+p_{6}=\frac{12}{21}
$$

Answer 11: There are $n$ students in a class. Let $A_{n}$ the event " at least 2 students have the same birthday from $n$ students ", so :

$$
\begin{aligned}
P\left(A_{n}\right) & =1-P\left(\overline{A_{n}}\right) \\
& =1-\frac{\operatorname{card}\left(\overline{A_{n}}\right)}{\operatorname{card}(\Omega)} \\
& =1-\frac{P_{365}^{n}}{365^{n}}
\end{aligned}
$$

If $n=2: P\left(A_{n=2}\right)=1-\frac{365 \times 364}{365^{2}}=0.0027$
Question : Find $n$, such as $P\left(A_{n}\right)=0.5$.
Exercise 12: We have 20 problems. The student can solve 11 of them. The instructor selects 5 questions randomly.

1. Let $A$ the event " the student can solve all five problems on the exam", so :

$$
P(A)=\frac{\operatorname{card}(A)}{\operatorname{Card}(\Omega)}=\frac{C_{11}^{5}}{C_{20}^{5}}=\ldots \ldots
$$

2. Let $B$ the event " the student can solve exactly one problem", so :

$$
P(B)=\frac{\operatorname{card}(A)}{\operatorname{Card}(\Omega)}=\frac{C_{11}^{1} \times C_{9}^{4}}{C_{20}^{5}}=\ldots \ldots
$$

3. Let $C$ the event " the student can solve at least two problems", so :

$$
P(C)=1-P(\bar{C})=1-\frac{\operatorname{card}(\bar{C})}{\operatorname{Card}(\Omega)}=1-\frac{C_{9}^{5}+C_{11}^{1} \times C_{9}^{4}}{C_{20}^{5}}=\ldots \ldots
$$

4. Let $D$ the event " the student can solve at most one problem", so :

$$
P(D)=P(\bar{C})=\ldots \ldots
$$

